Parameter, Input and State Estimation for Linear Structural Dynamics using parametric Model Order Reduction and Augmented Kalman Filtering

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Abstract

Tracking the evolution in time of parameters, input and states of a structural dynamic system is often difficult, since their direct measurement can be problematic or even impossible. It is of great interest to estimate these quantities based on output-only data from a limited set of sensors. This work proposes an estimation technique for states, inputs and material parameters for structural dynamics models based on an Augmented Extended Kalman Filter. A parametric Model Order Reduction technique is proposed to construct a Reduced Order Model which maintains an explicit dependency on material parameters, enabling the parameter estimation thanks to a low computational cost and an efficient derivation of the linearized system. The choice of sensor configurations that ensure the observability of unknown quantities is discussed as well. The proposed methodology shows highly promising results and could be employed for model refinement or condition monitoring. The methodology is validated both numerically and experimentally, using data acquired on a scaled wind turbine blade, with errors on the estimated parameters lower than 3.5% with respect to experimentally identified parameter values.

Keywords: Structural dynamics, Parameter-input-state estimation, Parametric model order reduction, Augmented extended Kalman filter, Parameter identification

1 1. Introduction

Parameter identification [1, 2] and tracking [3, 4] have been among the main research 2 interests in structural dynamics over the last years. The system's states are also com-3 monly tracked to have a complete knowledge about the system evolution in time. In most 4 cases, parameters and states are difficult to directly measure. The main limiting factors 5 are, among others, the high cost of the required sensors and mounting inaccessibility, 6 potential interference with operations or, at times, appropriate sensors might not even exist. A common approach to overcome these issues is employing indirect estimation 8 methods, that are typically based on the measurement of a set of responses of the system Preprint submitted to Mechanical Systems and Signal Processing September 30, 2022 to known inputs. Often however, measuring these inputs is also not straightforward.
 Therefore, ideally an output-only procedure should be employed, where the estimation
 of parameters and states is coupled with the concurrent estimation of inputs.

A well known estimation approach, originally developed for state estimation, is the 13 Kalman Filter (KF) [5, 6], which combines a numerical model of the system and measured 14 sensor data by minimizing the trace of the state error covariance matrix. For non-15 linear systems, several variations have been proposed in literature, such as e.g. the 16 Extended Kalman Filter (EKF) [6] and the Sigma-Point Kalman Filters (SPKF) [7]. In 17 18 order to simultaneously estimate inputs and parameters, common approaches are Dual Kalman Filters (DKF) [8, 9] or state augmentation [10]. DKF consists of two separate 19 filters estimating the states and the unknown quantities in parallel. State augmentation 20 consists instead of extending, i.e. augmenting, the system state vector with the unknown 21 quantities such as e.g. inputs or parameters. This approach allows to keep into account 22 all coupling effects between the estimated quantities, given that the augmented system 23 is observable [11]. Furthermore, it generally results in a gain in computational time. A 24 zeroth order random walk model is often used to model the augmented states dynamics 25 [12, 13]. By employing one of these two techniques, several authors jointly estimated 26 states and inputs [13, 14, 15], states and parameters [16, 17, 18, 9] or the three together 27 [12, 19, 20, 21].28

For structural dynamics applications, KF approaches typically use Finite Element 29 (FE) models, that can have a large number of Degrees of Freedom (DOFs); i.e. in 30 the order of hundreds of thousands up to millions. The direct use of these models 31 can result in an unfeasible computational load. Their dimensions are usually reduced 32 by employing projection-based Model Order Reduction (MOR) techniques [22, 23, 24]. 33 Here, the governing equations of the Full Order Model (FOM) are projected onto a 34 lower dimensional subspace of the solution space, the so-called reduction space, which is 35 spanned by a reduction basis. This effectively reduces both the number of equations and 36 DOFs, resulting in a so-called Reduced Order Model (ROM) with a decreased associated 37 computational load. 38

When using a KF for parameter estimation, the parameter values are updated at 39 each time step, requiring the model to be updated accordingly. Since in general ROMs 40 do not retain any explicit parametric dependency, it is necessary to recreate the ROM at 41 each step, which could result in a loss of most of the computational efficiency granted by 42 the use of a ROM in the first place. This explicit dependency can however be retained 43 by applying a parametric Model Order Reduction (pMOR) technique. A comprehensive 44 overview on projection-based pMOR techniques can be found in the work by Benner et 45 al. [25]. The usage of pMOR for model identification together with optimization methods 46 has been proposed in literature [26, 27]. Employing a parametric Reduced Order Model 47 (pROM) inside a KF-based parameter estimation allows to efficiently update the model 48 for each new estimated parameter value, resulting in an overall computationally efficient 49 procedure. This approach has been proposed by Naets et al. [12] for the joint estimation 50 of states, parameters and input in linear structural dynamics applications. Here, the 51 employed pMOR scheme consists of the interpolation of ROMs which have been calcu-52 lated for different sampled parameter values, which can be seen as a *local* approach [25]. 53 By not assuming any particular form for the parametric dependency of the full model, 54 this approach allows to have a general representation for a large variety of parameters. 55 This causes a loss of information about the kind of parameters considered, and the form 56

⁵⁷ of parametric dependency is not fully exploited. The downside of the local approach is ⁵⁸ that a ROM and its reduction basis have to be calculated and stored for each parame-⁵⁹ ter sample, potentially resulting in a large computational overhead. The interpolation ⁶⁰ of the ROMs and reduction bases introduces further computational bottlenecks. In or-⁶¹ der to guarantee the consistency of the different reduced sets of DOFs, an intermediate ⁶² projection to a high-dimensional space is typically required.

This work proposes an augmented EKF framework which uses a pMOR approach that 63 exploits models with an affine parameter dependency and uses a constant global reduction 64 65 basis [25]. More specifically, linear structural FE models show an affine dependency on material parameters. In this case, a global reduction basis, i.e. not depending on 66 parameters, can be defined and used to reduce the so-called affine components in an 67 off-line phase. The evaluation of the parametric model becomes then straightforward in 68 the on-line phase [28]. Only one global basis and a small set of matrices - the affine 69 components - thus have to be stored, leading to significant gains with respect to storage 70 and a positive impact on the computational performances. The proposed framework 71 enables the efficient joint estimation of states, inputs and parameters. Furthermore, the 72 proposed pMOR approach preserves the structure of the FOM parametric dependency, 73 allowing to distinguish between different types of parameters (e.g. stiffness-related and 74 mass-related) and consequently to analyze their observability individually. 75

The paper is structured as follows: in section 2 the affine representation for the structural FE system matrices is derived, and the proposed pMOR scheme is described and validated. In section 3 the EKF is described and in section 4 the influence of the type of measurements and estimated quantities on the system observability is highlighted. An extensive numerical validation of the estimation procedure is presented in section 5. An experimental validation of the proposed approach is presented in section 6. Finally, conclusions are given in section 7.

83 2. Parametric Model Order Reduction for Finite Element structural models

For a parametric linear structural FE model, the semi-discrete (continuous in time, discrete in space) Equations Of Motion (EOM) of the FOM can be written as:

$\mathbf{M}(\mathbf{p})\ddot{\mathbf{z}} + \mathbf{C}(\mathbf{p})\dot{\mathbf{z}} + \mathbf{K}(\mathbf{p})\mathbf{z} = \mathbf{S}\mathbf{u}$ (1)

where $\mathbf{z} \in \mathbb{R}^{n_{dof}}$ are the nodal DOFs, M, C and K are respectively the stiffness, 86 damping and mass matrices, \mathbf{S} is the input shape matrix that distributes the external 87 forces $\mathbf{u} \in \mathbb{R}^{n_f}$ on the DOFs of the system and $\mathbf{p} \in \mathbb{R}^{n_p}$ are the parameters of the 88 system. In this work the parameters of interest are material properties such as Young's 89 modulus (E), density (ρ) and Poisson's ratio (ν) for isotropic materials. Each of the 90 EOM components is dependent on time, but has been omitted from notation for clarity 91 purposes. The EOM are generally derived by assuming constant mass for the system, 92 while if the density is time-variant an additional related term appears in the equations 93 [29]. The assumption made in this work is that density varies slowly in time, a realistic 94 behavior for system identification or monitoring applications, so that the additional term 95 is assumed to be negligible. 96

97 2.1. Affine Representation

The system matrices should exhibit an affine dependency on the considered parameters to guarantee the efficiency of the global basis pMOR approach. The affine representation for a generic matrix $\mathbf{X}(\mathbf{p})$ can be stated as:

$$\mathbf{X}(\mathbf{p}) = \mathbf{X}_0 + \sum_i \mathbf{X}_i f_i(\mathbf{p})$$
(2)

where \mathbf{X}_0 is the constant term, \mathbf{X}_i are the affine components and $f_i(\mathbf{p})$ are the affine functions.

In the following sections, the affine representation of the matrices of the system is explicitly described. The FE formulation employed in this section is based on [30].

¹⁰⁵ Stiffness Matrix parameterization

The constitutive relationship for an isotropic material links the stress vector $\boldsymbol{\sigma} \in \mathbb{R}^6$ and strain vector $\boldsymbol{\epsilon} \in \mathbb{R}^6$ through the constitutive matrix $\mathbf{E} \in \mathbb{R}^{6 \times 6}$. This matrix has an affine relationship with the Young's modulus E and Poisson coefficient ν of the material:

$$\mathbf{E} = \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} \end{bmatrix} \frac{E\nu}{(1+\nu)(1-2\nu)} + \begin{bmatrix} 2\mathbf{I} & \mathbf{0} \\ \mathbf{0} & \mathbf{I} \end{bmatrix} \frac{E}{2(1+\nu)} = \mathbf{E}^{\lambda}\lambda + \mathbf{E}^{\mu}\mu$$
(3)

where λ and μ are the Lamé parameters.

In the former equation and in the following **I**, **0** and **1** represent respectively the identity matrix, the zero matrix and the all-ones matrix of appropriate dimensions.

If the material properties are assumed constant over the element volume, the same parameter dependence as in Equation 3 applies for the element stiffness matrix \mathbf{K}_e :

$$\mathbf{K}_e = \mathbf{K}_e^\lambda \lambda + \mathbf{K}_e^\mu \mu \tag{4}$$

The corresponding assembled system stiffness matrix \mathbf{K} retains the parametric relationship.

From Equation 4 it is clear that the two Lamé parameters themselves are the needed affine functions if the system is parameterized for both E and ν . It can be shown in an equivalent way that E is itself the affine function needed if ν is not considered in the parameterization. The final affine representation of the parametric stiffness matrix for a model containing multiple materials is then:

$$\mathbf{K}(\mathbf{p}) = \mathbf{K}_0 + \sum_i [\mathbf{K}_i^{\lambda} \lambda_i + \mathbf{K}_i^{\mu} \mu_i] + \sum_j \mathbf{K}_j^E E_j$$
(5)

where $\lambda_i = \lambda(E_i, \nu_i)$ and $\mu_i = \mu(E_i, \nu_i)$. The materials represented by an index i are parameterized for both E and ν , while the ones represented by an index j are only parameterized for E. \mathbf{K}_0 is a constant term accounting for possible terms that do not depend on parameters (if for example some of the materials in the model are not parameterized). The approach could be easily generalized to take into account localized stiffness terms related to lumped spring elements.

127 Mass Matrix parameterization

The consistent mass matrix of a single element, assuming constant material parameters over the volume, can be represented with an affine dependency on the material density ρ as:

$$\mathbf{M}_e = \mathbf{M}^{\rho} \rho \tag{6}$$

Following the same considerations made for the stiffness matrix, the affine representation of the parameterized mass matrix for a model containing multiple materials is:

$$\mathbf{M}(\mathbf{p}) = \mathbf{M}_0 + \sum_i \mathbf{M}_i^{\rho} \rho_i \tag{7}$$

where \mathbf{M}_0 is a constant term accounting for eventual terms that do not depend on parameters (such as non-parameterized lumped masses or materials).

¹³⁶ Damping Matrix parameterization

A typical and well known approach to model damping in structural applications is the proportional Rayleigh damping model. This consists in defining the damping matrix as a linear combination of mass and stiffness matrices as:

$$\mathbf{C}(\mathbf{p}) = \alpha \mathbf{M}(\mathbf{p}) + \beta \mathbf{K}(\mathbf{p}).$$
(8)

where α and β are the Rayleigh damping parameters.

This model has the advantage of retaining the parametric dependencies for the damp ing matrix.

¹⁴³ Affine Component Extraction

Starting from a set of matrices calculated at n_s different parameter values, the socalled sampling points, a least squares problem is solved to calculate the affine components. For a generic matrix, given the samples $[\mathbf{p}^1...\mathbf{p}^{n_s}]$ and the corresponding matrices $[\mathbf{X}^1...\mathbf{X}^{n_s}]$, the affine components are calculated by solving for each element $X_{(a,b)}$ of the matrices:

$$\min_{X_{0(a,b)},\dots,X_{i(a,b)},\dots} \sum_{h=1}^{n_s} \left\| X_{(a,b)}^h - X_{0(a,b)} - \sum_i X_{i(a,b)} f_i(\mathbf{p}^h) \right\| \tag{9}$$

where *i* represents the affine component index, as in Equation 2. In the former equation and in the rest of the paper, $\|\cdot\|$ denotes the L^2 norm. This procedure requires the sampling of the FOM a number of times at least equal to the number of components to be identified. An n_s lower than the number of unknowns in the least squares problem would in fact results in an underdetermined problem. The full matrices can be stored and used for identification during the sampling phase of the pROM creation (described below) at a minimum added computational cost.

156 2.2. pMOR Scheme

The DOFs vector \mathbf{z} is approximated by projecting it onto a lower-dimension subspace spanned by the global basis $\Psi \in \mathbb{R}^{n_{dof} \times n_{red}}$ that, given the pMOR scheme employed, is constant and does not depend on time or on the value of the parameters of interest:

$$\mathbf{z} \approx \mathbf{\Psi} \mathbf{q}$$
 (10)

160 where $\mathbf{q} \in \mathbb{R}^{n_{red}}$ is the reduced set of DOFs.

¹⁶¹ The Galerkin projection of the EOM onto the reduced space yields the reduced EOM:

$$\mathbf{M}^{r}(\mathbf{p})\ddot{\mathbf{q}} + \mathbf{C}^{r}(\mathbf{p})\dot{\mathbf{q}} + \mathbf{K}^{r}(\mathbf{p})\mathbf{q} = \mathbf{S}^{r}\mathbf{u}$$
(11)

¹⁶² where
$$\mathbf{M}^{r}(\mathbf{p}) = \boldsymbol{\Psi}^{T} \mathbf{M}(\mathbf{p}) \boldsymbol{\Psi}, \ \mathbf{C}^{r}(\mathbf{p}) = \boldsymbol{\Psi}^{T} \mathbf{C}(\mathbf{p}) \boldsymbol{\Psi}, \ \mathbf{K}^{r}(\mathbf{p}) = \boldsymbol{\Psi}^{T} \mathbf{K}(\mathbf{p}) \boldsymbol{\Psi} \text{ and } \mathbf{S}^{r} =$$

¹⁶³ $\boldsymbol{\Psi}^{T} \mathbf{S}.$

¹⁶⁴ Using Equation 5 and Equation 7, the parametric reduced matrices become:

$$\mathbf{K}^{r}(\mathbf{p}) = \mathbf{K}_{0}^{r} + \sum_{i} [\mathbf{K}_{i}^{r\lambda}\lambda(E_{i},\nu_{i}) + \mathbf{K}_{i}^{r\mu}\mu(E_{i},\nu_{i})] + \sum_{j} \mathbf{K}_{j}^{rE}E_{j}$$
(12)

$$\mathbf{M}^{r}(\mathbf{p}) = \mathbf{M}_{0}^{r} + \sum_{i} \mathbf{M}_{i}^{r\rho} \rho_{i}$$
(13)

¹⁶⁵ And the reduced damping matrix retains the proportional relationship:

$$\mathbf{C}^{r}(\mathbf{p}) = \alpha \mathbf{M}^{r}(\mathbf{p}) + \beta \mathbf{K}^{r}(\mathbf{p})$$
(14)

166 Reduction Basis Selection

In this work the global reduction basis Ψ is constructed by concatenating local bases Ψ_i , corresponding to different parameter samples \mathbf{p}_i , and carrying out a subsequent Singular Value Decomposition (SVD) [24].

The reduction basis is representative of the system for a set of configurations that 170 correspond to specific parameters values. This poses the challenge of identifying a global 171 reduction space that properly represents the system's behavior over the range of pa-172 rameter values of interest. This is particularly true for a modal basis, which contains 173 a subset of the eigenmodes of the system, since the eigenmode shapes can vary greatly 174 for different values of the parameters. The final reduction basis thus has to accurately 175 represent possible deformations across the parameter range in its entirety while having 176 a reasonably small dimension. The steps required for the basis selection are shown in 177 Figure 1, and are further described below. 178



Figure 1: Basis selection workflow

¹⁷⁹ Parameter Space Definition and Sampling

The parameter space Ω is defined by setting upper and lower limits for each parameter based on the application of interest: $\Omega = \{p_i \in \mathbb{R}, p_i^{min} < p_i < p_i^{max} \forall p_i\}.$

Once defined, the parameter space has to be sampled in order to get a set of config-182 urations that is representative of the entire space. Several sampling methods have been 183 proposed in literature for pMOR. Benner et al. [25] state that, for a small or moderate 184 number of parameters $(n_p < 10)$, a Latin Hypercube Sampling (LHS) method is effective. 185 Several methods have been tested for the application at hand (e.g. uniform sampling, 186 LHS) and LHS proved to be the most effective. Therefore this method is adopted in this 187 work. In case of a large parameter space, for which the LHS would require a high number 188 of samples, other techniques (such as e.g. greedy sampling [31]) could be adopted. 189

¹⁹⁰ Local Bases and Matrices Generation

For each parameter sample \mathbf{p}_i the full system, consisting of the matrices \mathbf{M}_i and \mathbf{K}_i , is assembled. These FOM matrices are used to extract the corresponding local basis. This basis consists of a set of eigenmodes $\Psi_{nm,i}$ concatenated with a set of static modes $\Psi_{sm,i}$ [13, 23], so that $\Psi_i = \Psi(\mathbf{p}_i) = [\Psi_{nm,i}\Psi_{sm,i}]$. The FOM matrices are stored for the affine component identification as explained in section 2.1.

¹⁹⁶ Global Reduction Basis Construction

¹⁹⁷ The global basis is constructed by concatenating the local bases:

$$\Psi_{global} = [\Psi_1 ... \Psi_{n_s}] \tag{15}$$

In general, this matrix of local bases can have linearly dependent columns. In order to remove these linear dependencies, an SVD is performed:

$$\Psi_{alobal} = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T \tag{16}$$

The Left Singular Vectors (LSVs) U represent an orthonormal set of vectors that, 200 spanning the range of Ψ_{qlobal} [24], is suitable to be used as a reduction basis. In order to 201 only keep the vectors that contain most of the information and therefore obtain a global 202 basis of small enough dimension, a subset of the LSVs is used as the global truncated 203 basis Ψ . This truncated LSVs basis exhibits the property of being the optimal low-rank 204 approximation for the full basis [24]. The choice of how many LSVs to retain is done 205 through an iterative procedure that subsequently increases the number of retained basis 206 vectors and compares the Frequency Response Functions (FRFs) of the pROM over a 207 set of n_{in} input DOFs and n_{out} output DOFs and in a specific frequency range until the 208 relative error becomes lower than a user-defined threshold. 209

Let $H_{ij}(\omega, \Psi, \mathbf{p})$ be the FRF for the input DOF *i* and the output DOF *j*, using the PROM created using the basis Ψ evaluated for the parameter values \mathbf{p} . Furthermore, the frequency range is discretized as $[\omega_1, \ldots, \omega_{n_\omega}]$. At each iteration step *k*, the relative error between the FRFs calculated using the current reduction basis Ψ_k and the one relative to the previous step Ψ_{k-1} is calculated as:

$$e_{k}(\mathbf{p}) = \sum_{i=1}^{n_{in}} \sum_{j=1}^{n_{out}} \left(\frac{1}{n_{\omega}} \sum_{l=1}^{n_{\omega}} \frac{\|(\|H_{ij}(\omega_{l}, \Psi_{k-1}, \mathbf{p})\| - \|H_{ij}(\omega_{l}, \Psi_{k}, \mathbf{p})\|)\|}{\|H_{ij}(\omega_{l}, \Psi_{k-1}, \mathbf{p})\|} \right)$$
(17)

This metric allows to achieve a good accuracy of the pROM in the frequency range of interest while limiting the number of evaluations needed for the FRFs to only the set of input/output DOFs. Furthermore, the metric also allows to compare both the shape of the responses and the amplitude. The error can be evaluated on a single parameter configuration of interest or averaged over a set of configurations. In this work, the error is calculated for the parameter values at the center of the sampled range.

221 2.3. pMOR Numerical Validation

A numerical validation of the proposed pMOR scheme is shown against a standard ROM and is presented below.

The system considered for the validation is a scaled wind turbine blade present at the testing facilities of Siemens Industry Software NV in Leuven (Belgium). The blade consists of Titanium Gr23 and has been produced using additive manufacturing followed by heat treatments. It is fixed to a concrete base via a set of bolted connections to its flange.

The linear FE model of the blade used in this work (Figure 2a) mainly consists of linear hexahedral elements and a minority of linear pentahedral elements, for a total of 65264 nodes and 391584 DOFs. Rigid elements connect the entire face of the flange to four spring elements that model the compliance of the bolts, as shown in Figure 2b.



Figure 2: Finite element model of the blade

The pROM has been created considering the Young's modulus and the density as parameters. This choice aims at giving a smaller set of representative parameters that can be easily represented in results.

A commercial FE solver [32] has been used to extract the full model information 236 needed for the pROM generation. The least-squares procedure identification described 237 in section 2 has been applied for the affine component identification. The settings used 238 for the creation of the pROM can be seen in Table 1. The number of eigenvectors 239 of the structure with an associated eigenfrequency in the range of interest is higher 240 for configurations with low stiffness and high density. The size of the local basis has 241 therefore been chosen to contain the entire set of modes even for these extreme cases. 242 Figure 3 shows the convergence of the relative error in the basis size selection procedure, 243 in which the basis starts with a dimension of 24 and is increased by 12 basis vectors at 244 each iteration. The relative error decreases monotonically to reach a value beyond the 245 threshold in 3 iterations. The final dimension of the pROM is 60, with a reduction factor 246 of more than 6000 with respect to the full model. 247

$\mathbf{size} \mathbf{of} \boldsymbol{\Psi}_i$	24 (15 eigenmodes + 9 static modes)
number of samples	5
relative error threshold	10^{-3}
frequency range	0-500Hz
number of nodes	65264
DOFs of the full model	391584
$\mathbf{size ~of}~ \boldsymbol{\Psi}_{global}$	60

Table 1: Blade's pROM information



Figure 3: Relative error for the global basis selection iterative procedure of the blade's pROM. The basis size starts at 24 and the first relative error is calculated at the second step where the size increased to 36.

248 Frequency Domain Analysis

In order to check that the pROM properly interpolates the behavior of the model 249 in the frequency range and parameter space considered, the FRFs can be calculated 250 and compared for a set of input/output DOFs of interest with the ones of the full FE 251 models and non-parametric ROMs created for a specific set of parameters. The results in 252 Figure 4 show the FRFs of a FE model, a standard, non-parametric ROM and the pROM 253 for parameter values not used in the pROM generation, namely the nominal values as well 254 as two extremes of the parameter range, for input and output on the force application 255 point at the bottom face of the blade. A good correspondence can be observed. 256



Figure 4: FRFs comparison on 3 parameter sets.

In Figure 5 the average error (calculated as $\frac{|(|FRF_{pROM}| - |FRF_{ROM}|)|}{|FRF_{ROM}|}$) on the FRFs between ROM and pROM is plotted over a parameter range larger than the one used for generating the pROM, corresponding to 11 × 11 samples. This shows as, in the original range, the interpolation by the pROM is accurate and provides good results even far from sampling points. Outside of the range the error is still low, but it shows a tendency to grow in the low stiffness zone where the frequency range includes a higher number of eigenfrequencies associated with eigenmodes of the structure.



Figure 5: Relative error between ROM and pROM FRFs over the parameter space sampled on a 11x11 grid. The grey box and the red stars are respectively the parameter space and the sampled points used for the pROM generation.

264 Time Domain Analysis

Since instabilities can arise when employing a ROM for time domain simulations if not 265 handled properly [33, 34], the comparison of FRFs might not be sufficient to properly 266 asses the quality of the proposed pMOR technique. Therefore, a comparison is also 267 carried out in the time domain. The results are shown in Figure 6, where a continuous 268 random signal with a frequency content in the range of 0-500Hz is applied to the pROM. 269 The acceleration is calculated at the same location as the input force, while the strain 270 is calculated at a location on the top face in the direction of the axis of the blade. A 271 first-order implicit Euler time integration scheme was used for the forward simulation 272 presented in this section with an integration step h = 0.1ms. A good correspondence 273 between the ROM and pROM can again be observed. 274



Figure 6: Comparison of time-domain simulations for a continuous random input. The right figures show a zoom of the 1.8s to 2s range.

3. Augmented Extended Kalman Filter for joint state/input/parameter es timation

In order to use the pROM within a Kalman Filter scheme, the reduced EOM are transformed from a second order to a first order form via a state space representation [13]. State augmentation is employed for the estimation of the inputs and parameters jointly with the states of the system. The augmented state is defined by extending the regular system states with the unknown quantities as $\mathbf{x}^* = \left[\mathbf{q}^T \, \dot{\mathbf{q}}^T \, \mathbf{u}^{unT} \, \mathbf{p}^T\right]^T \in \mathbb{R}^{2n_{red}+n_{un}+n_p}$. Here the input is split into the unknown inputs $\mathbf{u}^{un} \in \mathbb{R}^{n_{un}}$ and the known inputs $\mathbf{u}^{kn} \in \mathbb{R}^{n_{kn}}$, with a corresponding split in \mathbf{S}^r , giving $\mathbf{S}_{kn}^r \in \mathbb{R}^{n_{red} \times n_{kn}}$ and $\mathbf{S}_{un}^r \in \mathbb{R}^{n_{red} \times n_{un}}$.

The time evolution of inputs and parameters is represented using a zeroth order random walk model [10, 13, 12] where each augmented state is considered as constant with the addition of white Gaussian noise:

$$\dot{\mathbf{u}}^{un}(t) = \mathbf{0} + \mathbf{r}_u(t) \tag{18}$$

$$\dot{\mathbf{p}}(t) = \mathbf{0} + \mathbf{r}_p(t) \tag{19}$$

²⁸⁸ The augmented system can then be represented as:

$$\begin{cases} \dot{\mathbf{x}}^{*}(t) = \mathbf{A}^{*}(\mathbf{x}^{*})\mathbf{x}^{*}(t) + \mathbf{B}^{*}(\mathbf{x}^{*})\mathbf{u}^{kn}(t) \\ \mathbf{y}(t) = \mathbf{H}^{*}(\mathbf{x}^{*})\mathbf{x}^{*}(t) + \mathbf{D}^{*}(\mathbf{x}^{*})\mathbf{u}^{kn}(t) \end{cases}$$
(20)

in which the vector $\mathbf{y} \in \mathbb{R}^{n_m}$ contains the outputs from measurements, and the augmented state and input matrices are defined as:

$$\mathbf{A}^{*}(\mathbf{x}^{*}) = \begin{bmatrix} \mathbf{0} & \mathbf{I} & \mathbf{0} & \mathbf{0} \\ -\mathbf{M}^{r^{-1}}\mathbf{K}^{r} & -\mathbf{M}^{r^{-1}}\mathbf{C}^{r} & \mathbf{M}^{r^{-1}}\mathbf{S}_{un}^{r} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \end{bmatrix}$$
(21)

$$\mathbf{B}^{*}(\mathbf{x}^{*}) = \begin{bmatrix} \mathbf{0} \\ \mathbf{M}^{r^{-1}} \mathbf{S}_{kn}^{r} \\ \mathbf{0} \\ \mathbf{0} \end{bmatrix}$$
(22)

The output and feedthrough matrices $\mathbf{H}(\mathbf{p})$ and $\mathbf{D}(\mathbf{p})$ depend on the quantity measured and are explicitly defined in section 4. Note that here and in the following equations the dependency of \mathbf{K}^r , \mathbf{C}^r and \mathbf{M}^r on the parameters is not written explicitly, if not needed, for notation clarity.

The augmented model has a nonlinear dependency on the parameters, resulting in 295 general in a nonlinear dependency on the augmented state vector. Because of this, it 296 is necessary to employ a non-linear KF approach. Withouth loss of generality with 297 respect to other non-linear KF approaches, the discrete version of the EKF is employed 298 in this work. The EKF has a similar structure with respect to the one of the linear 299 KF, but uses a system linearization to propagate the state error covariance and calculate 300 the Kalman gain. Given the affine dependency of the reduced system matrices on the 301 parameters employed in the pMOR, the linearized system is expected to be smooth 302 and continuous. The next sections describe the time discretization of the system, the 303 linearization procedure, as well as the EKF algorithm. 304

305 3.1. Discretization and Linearization

The augmented state-space model in Equation 20 needs to be discretized in time by choosing an appropriate time integration scheme. This work adopts the method described by Risaliti et al. [15] for the discretization and linearization of the state-space equations, in which the integration scheme of choice is the first order implicit Euler scheme. This choice allows a larger time step size compared to explicit time integrator schemes.

The state evolution in time for a time step h is discretized as:

$$\mathbf{x}_k^* = \mathbf{x}_{k-1}^* + h\dot{\mathbf{x}}_k^* \tag{23}$$

³¹² The implicit equation governing the system is then:

$$\mathbf{g}_{d}(\mathbf{x}_{k-1}^{*}, \mathbf{x}_{k}^{*}, \mathbf{u}_{k}^{kn}) = \mathbf{x}_{k-1}^{*} - \mathbf{x}_{k}^{*} + h\left(\mathbf{A}^{*}\left(\mathbf{x}_{k}^{*}\right)\mathbf{x}_{k}^{*} + \mathbf{B}^{*}\left(\mathbf{x}_{k}^{*}\right)\mathbf{u}_{k}^{kn}\right) = \mathbf{0}$$
(24)

This equation is solved for \mathbf{x}_k in the EKF prediction step to advance the value of the states from k - 1 to k.

The EKF linearization requires the extraction of the underlying explicit equation associated with Equation 24 as detailed in [15]. This explicit function exists locally if the corresponding implicit function is continuously differentiable. This property is guaranteed by the adopted affine parameter dependency. The underlying explicit function can be found using a Taylor series expansion around the linearization point $[\mathbf{x}_{k-1}^{*,0} \mathbf{x}_{k}^{*,0}]$, truncated at the first order:

$$\mathbf{g}_{d}(\mathbf{x}_{k-1}^{*}, \mathbf{x}_{k}^{*}, \mathbf{u}_{k}^{kn}) \approx \mathbf{g}_{d}^{0} + \frac{\partial \mathbf{g}_{d}}{\partial \mathbf{x}_{k}^{*}} \bigg|_{0} (\mathbf{x}_{k}^{*} - \mathbf{x}_{k}^{*,0}) + \frac{\partial \mathbf{g}_{d}}{\partial \mathbf{x}_{k-1}^{*}} \bigg|_{0} (\mathbf{x}_{k-1}^{*} - \mathbf{x}_{k-1}^{*,0}) \approx 0$$
(25)

where $\mathbf{g}_d^0 = \mathbf{g}_d(\mathbf{x}_{k-1}^{*,0}, \mathbf{x}_k^{*,0}, \mathbf{u}_k^{kn})$. Furthermore, \mathbf{x}_k^* can be explicitly defined as:

$$\mathbf{f}_{d}(\mathbf{x}_{k-1}^{*}, \mathbf{u}_{k}^{kn}) = \mathbf{x}_{k}^{*} = \mathbf{x}_{k}^{*,0} - \left(\frac{\partial \mathbf{g}_{d}}{\partial \mathbf{x}_{k}^{*}}\Big|_{0}\right)^{-1} \left[\mathbf{g}_{d}^{0} + \frac{\partial \mathbf{g}_{d}}{\partial \mathbf{x}_{k-1}^{*}}\Big|_{0} (\mathbf{x}_{k-1}^{*} - \mathbf{x}_{k-1}^{*,0})\right]$$
(26)

³²² from which the Jacobian matrix is defined as:

$$\mathbf{J} = \frac{\partial \mathbf{f}_d(\mathbf{x}_{k-1}^*, \mathbf{u}_k)}{\partial \mathbf{x}_{k-1}^*} = -\left(\frac{\partial \mathbf{g}_d}{\partial \mathbf{x}_k^*}\right|_0^{-1} \left(\frac{\partial \mathbf{g}_d}{\partial \mathbf{x}_{k-1}^*}\right|_0^{-1}$$
(27)

323 The two derivatives required for the evaluation of the Jacobian matrix are:

$$\frac{\partial \mathbf{g}_d}{\partial \mathbf{x}_k^*} = -\mathbf{I} + h \left[\frac{\partial \mathbf{A}^*(\mathbf{x}_k^*)}{\partial \mathbf{x}_k^*} \mathbf{x}_k^* + \mathbf{A}^*(\mathbf{x}_k^*) \right] + h \left[\frac{\partial \mathbf{B}^*(\mathbf{x}_k^*)}{\partial \mathbf{x}_k^*} \mathbf{u}_k^{kn} \right]$$
(28)

$$\frac{\partial \mathbf{g}_d}{\partial \mathbf{x}_{k-1}^*} = \mathbf{I} \tag{29}$$

The derivatives of the system matrices $\mathbf{A}^*(\mathbf{x}^*)$ and $\mathbf{B}^*(\mathbf{x}^*)$ are straightforward to compute given the affine dependency employed for the pMOR. Given Equation 21 and Equation 22, $\mathbf{A}^*(\mathbf{x}^*)$ and $\mathbf{B}^*(\mathbf{x}^*)$ only depend on \mathbf{p} , hence only their derivatives with respect to \mathbf{p} are non-zero. As $\mathbf{A}^*(\mathbf{x}^*)$ and $\mathbf{B}^*(\mathbf{x}^*)$ are the result of the multiplication of the $\mathbf{M}^{r^{-1}}(\mathbf{p})$, $\mathbf{C}^r(\mathbf{p})$ and $\mathbf{K}^r(\mathbf{p})$ matrices, the chain rule for derivation is applied to these terms to obtain the derivatives with respect to \mathbf{p} .

The derivative of the stiffness matrix $\mathbf{K}^{r}(\mathbf{p})$ depends on the type of parameter considered:

$$\frac{\partial \mathbf{K}^{r}(\mathbf{p})}{\partial E_{i}} = \mathbf{K}_{i}^{r\lambda} \frac{\partial \lambda}{\partial E} + \mathbf{K}_{i}^{r\mu} \frac{\partial \mu}{\partial E} = \mathbf{K}_{i}^{r\lambda} \frac{\nu}{(1+\nu)(1-2\nu)} + \mathbf{K}_{i}^{r\mu} \frac{1}{2(1+\nu)}$$
(30)

$$\frac{\partial \mathbf{K}^{r}(\mathbf{p})}{\partial \nu_{i}} = \mathbf{K}_{i}^{r\lambda} \frac{\partial \lambda}{\partial \nu} + \mathbf{K}_{i}^{r\mu} \frac{\partial \mu}{\partial \nu} = \mathbf{K}_{i}^{r\lambda} \frac{E(1+2\nu^{2})}{(1+\nu)^{2}(1-2\nu)^{2}} - \mathbf{K}_{i}^{r\mu} \frac{E}{2(1+\nu)^{2}}$$
(31)

$$\frac{\partial \mathbf{K}^{r}(\mathbf{p})}{\partial E_{j}} = \mathbf{K}_{j}^{E}$$
(32)

The derivative of the inverse mass matrix $\mathbf{M}^{r^{-1}}(\mathbf{p})$ is needed and can be calculated by using the property of inverse matrices derivation:

$$\frac{\partial \mathbf{M}^{r^{-1}}(\mathbf{p})}{\partial \rho_i} = -\mathbf{M}^{r^{-1}}(\mathbf{p}) \frac{\partial \mathbf{M}^r(\mathbf{p})}{\partial \rho_i} \mathbf{M}^{r^{-1}}(\mathbf{p}) \qquad \qquad \frac{\partial \mathbf{M}^r(\mathbf{p})}{\partial \rho_i} = \mathbf{M}_i^{r\rho} \qquad (33)$$

Given the proportional damping model used, the derivative of $\mathbf{C}^{r}(\mathbf{p})$ is a linear combination of the two calculated above.

336 3.2. Augmented Extended Kalman Filter

The discrete system equations of the system considering noise are:

$$\mathbf{g}_d(\mathbf{x}_{k-1}^*, \mathbf{x}_k^*, \mathbf{u}_k^{kn}) = w_k \tag{34}$$

$$\mathbf{y}_k = \mathbf{h}(\mathbf{x}_k^*, \mathbf{u}_k^{kn}) + v_k \tag{35}$$

where w_k and v_k are white uncorrelated Gaussian noise with the corresponding covariance matrices \mathbf{Q}_k and \mathbf{R}_k . The output Equation 35 is commonly referred to as the measurement equation in the context of Kalman filtering.

³⁴⁰ The discrete-time version of the augmented EKF consists of the following steps:

Prediction

$$\mathbf{g}_d(\mathbf{x}_{k-1}^{*+}, \mathbf{x}_k^{*-}, \mathbf{u}_k^{kn}) = 0 \tag{36}$$

$$\mathbf{J}_{k-1} = \frac{\partial \mathbf{f}_d(\mathbf{x}_{k-1}^*, \mathbf{u}_k^{kn})}{\partial \mathbf{x}_{k-1}^*} \bigg|_{\mathbf{x}_k^{*+}, \mathbf{x}_k^{*-}}$$
(37)

$$\mathbf{P}_{k}^{-} = \mathbf{J}_{k-1}\mathbf{P}_{k-1}^{+}\mathbf{J}_{k-1}^{T} + \mathbf{Q}_{k-1}$$
(38)

Correction

$$\mathbf{J}_{m,k} = \frac{\partial \mathbf{h}(\mathbf{x}_k^*, \mathbf{u}_k^{kn})}{\partial \mathbf{x}_k^*} \bigg|_{\mathbf{x}_k^{*-}}$$
(39)

$$\mathbf{K}_{k} = \mathbf{P}_{k}^{-} \mathbf{J}_{m,k}^{T} (\mathbf{J}_{m,k} \mathbf{P}_{k}^{-} \mathbf{J}_{m,k}^{T} + \mathbf{R}_{k})^{-1}$$
(40)

$$\mathbf{x}_{k}^{*+} = \mathbf{x}_{k}^{*-} + \mathbf{K}_{k}(\mathbf{y}_{k} - \mathbf{h}(\mathbf{x}_{k}^{*-}, \mathbf{u}_{k}^{kn}))$$
(41)

$$\mathbf{P}_{k}^{+} = (\mathbf{I} - \mathbf{K}_{k} \mathbf{J}_{m,k}) \mathbf{P}_{k}^{-} (\mathbf{I} - \mathbf{K}_{k} \mathbf{J}_{m,k})^{T} + \mathbf{K}_{k} \mathbf{R}_{k} \mathbf{K}_{k}^{T}$$
(42)

In the former equations, \mathbf{J}_{k-1} is the Jacobian of the system calculated on the a-posteriori estimate from the previous step. The linearized form of the measurement equation around the current estimated state is represented by $\mathbf{J}_{m,k}$ The term \mathbf{K}_k is the Kalman gain, and \mathbf{P}_k is the state error covariance matrix.

Of particular importance in the setup of the filter is the choice of the matrices \mathbf{Q} and \mathbf{R} , respectively representing the covariance of the states (also referred to as the plant noise covariance matrix) and of the measurements. They are commonly assumed as constant and set by trial and error procedure or by experience. In literature some rules of thumb for the choice have been proposed (e.g. [13]) or schemes to adapt the value in time (e.g. [35, 18, 36]). In particular, in this work the \mathbf{R} matrix is chosen to be a constant diagonal matrix containing the measurement noise covariance taken directly from the used physical sensor data. The plant noise covariance matrix \mathbf{Q} is also assumed to be constant and diagonal, as the noise is considered to be uncorrelated, and can be split to separate the contributions to the different augmented states as:

$$\mathbf{Q} = \begin{bmatrix} \mathbf{Q}_{q} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{Q}_{\dot{q}} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{Q}_{u} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{Q}_{p} \end{bmatrix}$$
(43)

Here, \mathbf{Q}_q and $\mathbf{Q}_{\dot{q}}$ represent the uncertainty on the model, while \mathbf{Q}_u and \mathbf{Q}_p represent the uncertainty on the unknown input and parameters, respectively.

The terms of \mathbf{Q}_u are set using the rule of thumb proposed in [13]. For the i-th unknown input u_i , the relative term on the diagonal of \mathbf{Q} is set as:

$$Q_{u_i} = \left(h\omega_{u_i}a_{u_i}\right)^2 \tag{44}$$

where ω_{u_i} and a_{u_i} are the expected values of frequency and amplitude of the force. As pointed out in [13] and [15], this method aims at keeping the rate of change of the unknown quantity in a range that allows to correctly track its variation while filtering noise. Following this reasoning, \mathbf{Q}_p is also set using the same rule of thumb:

$$Q_{p_i} = \left(h\delta_{p_i}a_{p_i}\right)^2 \tag{45}$$

In this case a_p is the initial tentative value for the parameter and δ_p is a scaling factor that helps in the setup of the unknown parameter covariance by decoupling it from the absolute value of the parameter itself as well as the integration scheme time step size. In this way, the influence of δ_p on the estimation results can be studied and applied to parameters with different order of magnitudes. The units of δ_p are equivalent to frequency.

The value of δ_p should typically be set to a low value, as the parameter value is usually constant or slowly varying. It should not be set excessively low, as this can slow down the convergence rate to as well as constrain the tracking of any change of the parameter value in time. Given the adopted zeroth order random walk model, a higher augmented state covariance value for the corresponding parameter state allows for a larger change per timestep, lowering the achieved filtering of noise from the estimated augmented states, which is present due to measurement noise propagation in the correction step of the EKF.

376 4. Observability analysis

In order for the EKF to be able to correctly estimate the unknown quantities of the system, a proper choice of the measurement set is needed [12, 14, 37]. This choice can be made by analyzing the system observability, which depends on the system-measurement combination $(\mathbf{A}^*, \mathbf{H}^*)$.

381 4.1. Continuous System Linearization

As noted in [12], observability analysis methods are traditionally developed for linear systems, but given the mild non-linearity of structural systems with respect to parameters it is still worth analyzing the properties of the linearized model around representative linearization points. Furthermore, the analysis is done on the time-continuous system for simplicity reasons, but all results translate directly to the discrete case. The continuous linearized model is represented by \mathbf{J}_c and \mathbf{J}_m , which are the derivatives of the state and measurement equations of the continuous-time model in Equation 20 respectively, with respect to the system state.

³⁹⁰ The Jacobian of the continuous state equation is defined as:

$$\mathbf{J}_{c} = \frac{\partial \dot{\mathbf{x}}^{*}}{\partial \mathbf{x}^{*}} = \mathbf{A}^{*}(\mathbf{x}^{*}) + \frac{\partial \mathbf{A}^{*}(\mathbf{x}^{*})}{\partial \mathbf{x}^{*}} \mathbf{x}^{*} + \frac{\partial \mathbf{B}^{*}(\mathbf{x}^{*})}{\partial \mathbf{x}^{*}} \mathbf{u}^{kn} = \begin{bmatrix} \mathbf{0} & \mathbf{I} & \mathbf{0} & \mathbf{0} \\ -\mathbf{M}^{r^{-1}} \mathbf{K}^{r} & -\mathbf{M}^{r^{-1}} \mathbf{C}^{r} & \mathbf{M}^{r^{-1}} \mathbf{S}_{un}^{r} & \frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \end{bmatrix}$$
(46)

391 where

$$\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}} = -\frac{\partial \mathbf{M}^{r^{-1}} \mathbf{K}^r}{\partial \mathbf{p}} \mathbf{q} - \frac{\partial \mathbf{M}^{r^{-1}} \mathbf{C}^r}{\partial \mathbf{p}} \dot{\mathbf{q}} + \frac{\partial \mathbf{M}^{r^{-1}}}{\partial \mathbf{p}} \mathbf{S}_{un}^r \mathbf{u}^{un} + \frac{\partial \mathbf{M}^{r^{-1}}}{\partial \mathbf{p}} \mathbf{S}_{kn}^r \mathbf{u}^{kn}$$
(47)

Some of the terms in Equation 46 and Equation 47 become zero for specific choices of parameter types or specific types of input. If only a set of stiffness parameters \mathbf{p}_s is estimated, $\frac{\partial \dot{\mathbf{q}}}{\partial \mathbf{p}}$ takes the form:

$$\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}_{s}} = -\mathbf{M}^{r^{-1}} \frac{\partial \mathbf{K}^{r}}{\partial \mathbf{p}_{s}} \left(\mathbf{q} + \beta \dot{\mathbf{q}}\right) = -\mathbf{M}^{r^{-1}} \frac{\partial \mathbf{K}^{r}}{\partial \mathbf{p}_{s}} \mathbf{K}^{r^{-1}} \left(\mathbf{S}_{un}^{r} \mathbf{u}^{un} + \mathbf{S}_{kn}^{r} \mathbf{u}^{kn} - \mathbf{M}^{r} \left(\ddot{\mathbf{q}} + \alpha \dot{\mathbf{q}}\right)\right)$$
(48)

If only a set of mass parameters \mathbf{p}_m is to be estimated, $\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}}$ takes the form:

$$\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}_m} = -\frac{\partial \mathbf{M}^{r^{-1}}}{\partial \mathbf{p}_m} \left(\mathbf{K}^r \left(\mathbf{q} + \beta \dot{\mathbf{q}} \right) - \mathbf{S}^r_{kn} \mathbf{u}^{kn} - \mathbf{S}^r_{un} \mathbf{u}^{un} \right) = -\mathbf{M}^{r^{-1}} \frac{\partial \mathbf{M}^r}{\partial \mathbf{p}_m} (\ddot{\mathbf{q}} + \alpha \dot{\mathbf{q}})$$
(49)

For a measurement in which the dependency on states and inputs can be written in the form:

$$\mathbf{h}(\mathbf{x}^*, \mathbf{u}^{kn}) = \mathbf{H}^*(\mathbf{x}^*)\mathbf{x}^* + \mathbf{D}^*(\mathbf{x}^*)\mathbf{u}^{kn}$$
(50)

as it is the case in this work and without loss of generality, the Jacobian results in:

$$\mathbf{J}_{m} = \frac{\partial \mathbf{h}(\mathbf{x}^{*}, \mathbf{u}^{kn})}{\partial \mathbf{x}^{*}} = \mathbf{H}^{*}(\mathbf{x}^{*}) + \frac{\partial \mathbf{H}^{*}(\mathbf{x}^{*})}{\partial \mathbf{x}^{*}} \mathbf{x}^{*} + \frac{\partial \mathbf{D}^{*}(\mathbf{x}^{*})}{\partial \mathbf{x}^{*}} \mathbf{u}^{kn}$$
(51)

The observability analysis requires the explicit definition of equations for the different kinds of measurements. In this work, strain and acceleration measurements are considered:

402 Strain Measurements

The strain measurement equation can be derived as the spatial derivative of the nodal displacements. Given the element shape functions, the strain measurement equation can then be stated as [15]:

$$\mathbf{h}_s(\mathbf{x}^*, \mathbf{u}^{kn}) = \mathbf{N}_s \mathbf{z} = \mathbf{N}_s \boldsymbol{\Psi} \mathbf{q}$$
(52)

where $\mathbf{N}_s \in \mathbb{R}^{n_{ms} \times n_{dof}}$ is a shape matrix relating the physical DOFs to the measured strain components in the material points of interest and n_{ms} is the number of strain measurements acquired. In this case:

$$\mathbf{H}_{s}^{*} = \begin{bmatrix} \mathbf{N}_{s} \Psi & \mathbf{0} & \mathbf{0} \end{bmatrix} \qquad \qquad \mathbf{D}_{s}^{*} = \mathbf{0} \tag{53}$$

406 and

$$\mathbf{J}_{m,s} = \mathbf{H}_s^* \tag{54}$$

407 Acceleration Measurements

The acceleration measurement is can be expressed as a linear combination of nodal accelerations, so that:

$$\mathbf{h}_{a}(\mathbf{x}^{*},\mathbf{u}^{kn}) = \mathbf{N}_{a}\ddot{\mathbf{z}} = \mathbf{N}_{a}\boldsymbol{\Psi}\ddot{\mathbf{q}} = \mathbf{N}_{a}\boldsymbol{\Psi}\mathbf{M}^{r^{-1}}\left(-\mathbf{C}^{r}\dot{\mathbf{q}} - \mathbf{K}^{r}\mathbf{q} + \mathbf{S}_{kn}^{r}\mathbf{u}^{kn} + \mathbf{S}_{un}^{r}\mathbf{u}^{un}\right)$$
(55)

where $\mathbf{N}_a \in \mathbb{R}^{n_{ma} \times n_{dof}}$ is a shape matrix relating the physical acceleration DOFs to the acceleration of the material points of interest and and n_{ma} is the number of acceleration measurements acquired. In this case:

$$\mathbf{H}_{a}^{*}(\mathbf{x}^{*}) = \mathbf{N}_{a} \boldsymbol{\Psi} \begin{bmatrix} -\mathbf{M}^{r^{-1}} \mathbf{K}^{r} & -\mathbf{M}^{r^{-1}} \mathbf{C}^{r} & \mathbf{M}^{r^{-1}} \mathbf{S}_{un}^{r} & \mathbf{0} \end{bmatrix}$$
(56)

413

$$\mathbf{D}_{a}^{*}(\mathbf{x}^{*}) = \mathbf{N}_{a} \boldsymbol{\Psi} \left[\mathbf{M}^{r^{-1}} \mathbf{S}_{kn}^{r} \right]$$
(57)

and, by considering the dependency of the reduced matrices with respect to the augmented states (the parameters in particular)

$$\mathbf{J}_{m,a} = \mathbf{N}_{a} \boldsymbol{\Psi} \begin{bmatrix} -\mathbf{M}^{r^{-1}} \mathbf{K}^{r} & -\mathbf{M}^{r^{-1}} \mathbf{C}^{r} & \mathbf{M}^{r^{-1}} \mathbf{S}_{un}^{r} & \frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}} \end{bmatrix}$$
(58)

416 4.2. PBH Test

⁴¹⁷ The Popov-Belevitch-Hautus (PBH) observability test [38, 39] states that a sufficient ⁴¹⁸ and necessary condition for the observability of the linearized system is that the matrix

$$\mathbf{PBH} = \begin{bmatrix} \mathbf{I}s - \mathbf{J}_c \\ \mathbf{J}_m \end{bmatrix}$$
(59)

⁴¹⁹ is of full rank for every value of $s \in \mathbb{C}$. The matrix $\mathbf{I}s - \mathbf{J}_c$ is guaranteed to be full rank ⁴²⁰ for all values of s but the eigenvalues of \mathbf{J}_c , which are the values on which the analysis ⁴²¹ focuses. In particular, the linearized matrix has a number of zero-valued rows at least ⁴²² equal to the number of augmented states, thus it has the same number of zero eigenvalues. ⁴²³ It is for this value of s that the observability is most critical; non-zero eigenvalues, while ⁴²⁴ being unobservable, remain stable and thus the system is at least detectable [14]. The ⁴²⁵ s = 0 case will then be analyzed in the rest of this section. The PBH matrix, for s = 0 and using strain and acceleration measurements, can be expressed as:

$$\mathbf{PBH} = \begin{bmatrix} \mathbf{0} & -\mathbf{I} & \mathbf{0} & \mathbf{0} \\ \mathbf{M}^{r^{-1}}\mathbf{K}^{r} & \mathbf{M}^{r^{-1}}\mathbf{C}^{r} & -\mathbf{M}^{r^{-1}}\mathbf{S}_{un}^{r} & -\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{N}_{s}\Psi & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ -\mathbf{N}_{a}\Psi\mathbf{M}^{r^{-1}}\mathbf{K}^{r} & -\mathbf{N}_{a}\Psi\mathbf{M}^{r^{-1}}\mathbf{C}^{r} & \mathbf{N}_{a}\Psi\mathbf{M}^{r^{-1}}\mathbf{S}_{un}^{r} & \mathbf{N}_{a}\Psi\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}} \end{bmatrix}$$
(60)

428 Some conclusions can be drawn directly:

In order to have full observability, a number of strain measurements at least equal to the total number of unknown quantities (inputs and parameters) should be employed. The rows relative to acceleration measurements in Equation 60 are in fact always a linear combination of the second block rows. For any application case then, the use of only acceleration measurements does not allow for an observable system. Acceleration measurements are still important since they add information useful to improve the performances of the estimation procedure

• Parameters related to mass cannot be observed in a static condition (for which $\frac{\partial \ddot{\mathbf{q}}}{\partial \mathbf{p}}$ is zero) because the last block column becomes zero. This follows the practical intuition that in a static case the inertial term is not excited and thus measurements carry no useful information.

Other conclusions can be made in particular cases for which the equations simplify
and it is possible to observe linear dependency of the columns of the PBH matrix. In
the next subsection the observability for the case of a single parametrized material is
discussed.

444 4.3. Observability Analysis for a Single Material

This section considers the case for which one material is parameterized for both Young's modulus and density. In this case the mass, stiffness and damping matrices assume the simple form:

$$\mathbf{K}^r = E \mathbf{K}^{rE} \tag{61}$$

$$\mathbf{M}^{r} = \rho \mathbf{M}^{r\rho} \tag{62}$$

$$\mathbf{C}^{r} = \alpha \rho \mathbf{M}^{r\rho} + \beta E \mathbf{K}^{rE} \tag{63}$$

⁴⁴⁸ This simplified case yields the following derivatives:

$$\frac{\partial \ddot{\mathbf{q}}}{\partial E} = -\frac{1}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{K}^{rE} \left(\mathbf{q} + \beta \dot{\mathbf{q}} \right) = \frac{1}{\rho E} \mathbf{M}^{r\rho^{-1}} \left(\mathbf{S}_{kn}^{r} \mathbf{u}^{kn} + \mathbf{S}_{un}^{r} \mathbf{u}^{un} \right) - \frac{1}{E} (\ddot{\mathbf{q}} + \alpha \dot{\mathbf{q}}) \quad (64)$$

$$\frac{\partial \ddot{\mathbf{q}}}{\partial \rho} = -\frac{E}{\rho^2} \mathbf{M}^{r\rho^{-1}} \left(\mathbf{K}^{rE} \left(\mathbf{q} + \beta \dot{\mathbf{q}} \right) - \mathbf{S}_{kn}^r \mathbf{u}^{kn} \right) = -\frac{1}{\rho} (\ddot{\mathbf{q}} + \alpha \dot{\mathbf{q}})$$
(65)

449 For s = 0, the PBH matrix is:

$$\mathbf{PBH} = \begin{bmatrix} \mathbf{0} & -\mathbf{I} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \frac{E}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{K}^{rE} & \left(\alpha + \frac{\beta E}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{K}^{rE} \right) & -\frac{1}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{S}_{un}^{r} & -\frac{\partial \ddot{\mathbf{q}}}{\partial E} & -\frac{\partial \ddot{\mathbf{q}}}{\partial \rho} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ \mathbf{N}_{s} \Psi & \mathbf{0} & \mathbf{0} & \mathbf{0} & \mathbf{0} \\ -\mathbf{N}_{a} \Psi \frac{E}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{K}^{rE} & -\mathbf{N}_{a} \Psi \left(\alpha + \frac{\beta E}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{K}^{rE} \right) & \mathbf{N}_{a} \Psi \frac{1}{\rho} \mathbf{M}^{r\rho^{-1}} \mathbf{S}_{un}^{r} & \mathbf{N}_{a} \Psi \frac{\partial \ddot{\mathbf{q}}}{\partial E} & \mathbf{N}_{a} \Psi \frac{\partial \ddot{\mathbf{q}}}{\partial \rho} \end{bmatrix}$$
(66)

450 It can be observed that:

- Given Equation 64 and Equation 65, it is impossible to simultaneously estimate Eand ρ if there is no input being applied to the system. This implies that under a free response condition the effect of mass and inertia is not distinguishable. The last two columns become in fact linearly dependent.
- It is impossible to simultaneously estimate inputs, E, ρ if there is not at least one non-zero known input.
- If E and the input are to be estimated, the system is not observable in static conditions. This implies that it is impossible to estimate both the Young's modulus and a static load, confirming the intuition that the system inversion in this case is not feasible.

461 5. Numerical Validation

This section presents a numerical validation of the main methodology proposed and presented in the previous chapters, i.e. the joint state/input/parameter estimation, using the blade FE model introduced in subsection 2.3. In a first part an Optimal Sensor Placement (OSP) procedure is employed to select the sensors to use for the estimation. Then a setup of the filter is executed using numerical data. The robustness and effectiveness of the chosen setup is numerically validated.

468 5.1. Sensor Selection

The physical blade is instrumented with strain and acceleration sensors spread over its entire length, as shown in Figure 7, therefore also virtual sensors are placed at the same locations. In particular, the sensors used are:

- 10 rectangular strain gage rosettes placed symmetrically on the top and bottom face at 5 different sections along the length of the blade (numbered from 0 to 4 in the remaining part of the paper).
- 475
 4 uniaxial strain gages placed symmetrically on the top and bottom face of sections
 2 and 3 along the axis of the blade.

• 10 triaxial accelerometers placed on the top face at 8 sections along the length of the blade (named from A to H in the remaining part of the paper). Their axes are aligned with the global ones.



Figure 7: Position of sensors and input points on the blade. The 3 sensors labeled are the rosettes of which the z sensors were selected by the OSP procedure.

Only a subset of sensors is used for the estimation, while the remaining ones are used 480 for validation purposes: the estimated values are compared with the simulated ones, for 481 the numerical applications, or with the measured ones, for experimental applications. 482 A minimal number of strain gages is required for observability reasons, as discussed in 483 section 4. The addition of accelerometers to this set could extend the overall bandwidth 484 of the filter. The frequency range investigated in this work is however low enough to 485 allow the use of only strain gages for the estimation, while the accelerations sensors will 486 therefore be used for validation. 487

The selection of strain gages to use for the estimation has been based on the OSP 488 strategy described in [37]. There the aim is to find the optimal location and orientation 489 of sensors in order to guarantee the best observability for unknown inputs. The strategy 490 starts by screening an initial sensor pool spread over the entire surface of the model, and 491 then subsequently removing the sensors that contribute the least to the observability 492 metric of interest. For the blade setup considered in this work, the OSP strategy has 493 been adapted to start from the pool of existing sensors, following the subsequent steps 494 as discussed in the reference. 495

The OSP procedure has been used to select a set of 3 strain gages, since that is the minimum required number for estimation of the two parameters and one input. The selected sensors are the ones oriented along the axis of the blade in the rosettes on the top face of sections 2 and 4 and on the bottom face of section 4, as shown on Figure 7. The OSP approach targets input estimation and therefore it is only used in this work as

477 478 479 a way to select the sensors to use, without any claim of global optimality. The extension
 of the OSP procedure to consider also observability of parameters is out of the scope of
 this work and is therefore not explored here.

504 5.2. Filter Setup

In the case of experimental data, the measurement noise covariance matrix **R** is set using values extracted for each physical sensor. If numerical measurements are used, the noise level is known and its covariance value can be directly used as the diagonal entries of **R**. The initial value for **P** is set by assuming a low initial error on the states and an initial error on the parameters in the same order of magnitude as the expected parameter values themselves to improve initial convergence.

It is common practice for AKF application to set the covariance of the plant to a low value to let the filter follow the numerical model predictions [13, 12]. In the following application $\mathbf{Q}_q = \epsilon \mathbf{I}$ and $\mathbf{Q}_{\dot{q}} = \epsilon \mathbf{I}$, where ϵ represents the machine precision.

Setting the zeroth order random walk noise covariance to a low value implies that 514 the associated augmented state is expected to remain constant or vary slowly. A large 515 value permits a large variation of the associated augmented state during each time step. 516 These covariance values depend on the assumed nature of the augmented states, i.e. 517 having a specific expected quasi-static or dynamic behavior. In the remaining part of 518 this section the effect of choosing different augmented state noise covariance values on the 519 corresponding estimation results is shown. The values are represented by ω_u or δ_p , which 520 are used to calculate the covariance values by applying Equation 44 or Equation 45. The 521 following estimation cases are based on numerical data in order to have full control and 522 assure that any variations in the estimated values are only due to the noise present in 523 the supplied numerical measurements. The measurements are generated by a forward 524 simulation using the same integration scheme as in subsection 2.3 and then zero-mean 525 white Gaussian noise is added. A 20 Hz sinusoidal load (Figure 8) is used for this analysis, 526 without loss of generality. 527



Figure 8: 20Hz sine input

The first case considered is the estimation of a constant Young's modulus. The estimated parameter value for different values of δ_E is shown in Figure 9. In order to evaluate the estimation performance, the average error is considered along with the maximum and minimum error that define the interval in which the estimated parameter evolves. Ideally the average value for the error should be close to zero and the maximum and minimum interval should be as narrow as possible. A large interval indicates a large

variation usually caused by noise. The desired behavior of the filter in this first case is 534 that the estimated parameter value converges and stays as constant as possible. All the 535 δ_E values allow for similar convergence behavior, with a similar and small average error 536 in the last part. It can also be seen that the evolution interval is narrow for δ_E values 537 below a certain threshold $(10^{-2}Hz)$ while for higher values it increases, showing more 538 measurement noise leaking into the estimate, as explained in subsection 3.2. It can be 539 concluded in this case that a covariance value as low as possible is desired, and that any 540 value below a certain threshold will allow for good results. 541



Figure 9: Estimated value of E with numerical data for sine load with different values of δ_E . A zoom on the last 0.4s is presented together with the estimation error on the same time interval.

A second case is considered where the parameter is degrading in time (Figure 10). A low δ_E value constrains the rate of change of the parameter, preventing it from properly tracking the real value. The minimum δ_E value that allows to properly track the evolution of the parameter for the presented case is 1Hz. By further increasing the δ_E value, the average error remains small, while the variation interval gets larger, showing that more measurement noise leaks into the estimated augmented states. In this case a larger δ_E value should be used.

These results indicate that there is no overall optimal covariance value setting and the setting thus has to be made according to the application case and the expected evolution of the augmented state over time. A representative virtual exercise could be carried out to select these settings. In a system identification case, where the model parameters are expected to be constant, δ_E can be set as low as possible. If an abrupt change in parameter values is expected instead, a high value should to be employed at the expense of a more noisy estimate (as explained in subsection 3.2).



Figure 10: Estimated value of degrading E with numerical data for sine load with different values of δ_E . A zoom on the last 0.4s is presented together with the estimation error on the same time interval.

A third case is considered where E and ρ are estimated simultaneously, as shown 556 in Figure 11. If the covariance values are not high enough, the estimated parameters 557 tend to converge much slower as compared to the single parameter case. Also here a co-558 variance value for which the average error stabilizes corresponds to 1Hz. The decreased 559 convergence rate can be explained by the fact that the dynamic stiffness of a linear struc-560 tural system tends to depend on the ratio between stiffness and mass parameters. Once 561 the filter has reached the correct stiffness/mass ratio (by means of the estimated parame-562 ters), the dynamic response of the system using these estimated parameter values is close 563 to the real response, causing the low convergence rate that thus has to be compensated 564 for with higher covariance values of the parameters. This is shown in Figure 12. Here, 565 the ratio of the two parameters converges faster than their single values. 566

When parameters and input are estimated simultaneously, their covariance settings 567 appears to be uncoupled for this specific application, since they only effect their relative 568 augmented state's convergence. As a consequence, the choice of covariance values for 569 parameters, as discussed above, also applies for this case. The results shown in Figure 13 570 are calculated for a fixed $\delta_E = 10^{-3}$ and varying ω_u . As discussed and shown in [12, 40, 571 41], a small lag is usually present in the estimated input with respect to the measured 572 one, even when the amplitude is correctly estimated. This lag can also be clearly seen 573 in the results of this application when the frequency of the input is set to the known 574 value of 20Hz. This represents a mismatch in the case of input only estimation. When 575 parameters are jointly estimated with the inputs, this phase shift is compensated for by 576 the filter through an estimated parameter value different from the real one. Increasing 577 the estimated input frequency value reduces the observed lag at the expense of more 578



⁵⁷⁹ noise. This in turn reduces the error on the parameter estimation, as can be clearly seen ⁵⁸⁰ in the plot of parameter error with respect to ω_u .

Figure 11: Estimated values of E and ρ with numerical data for sine load with different values of $\delta_p = \delta_E = \delta_{\rho}$. A zoom on the last 0.4s is presented together with the estimation error on the same time interval.



Figure 12: Value of the ratio between estimated E and ρ for the case shown in Figure 11.



Figure 13: Estimated values of E and inputs with numerical data for sine load with different values of ω_u and $\delta_E = 10^{-3} Hz$. A zoom on the last 0.2s is presented together with the estimation error on the same time interval.

581 5.3. Numerical Estimation Validation

A first validation of the estimation procedure is done using numerical data. This allows, by having full control over the system and the measured data, to asses the performance of the filter with the setup defined in the previous section. The numerical approach also allows to consider the case of a parameter degrading with time, for which no experimental data is available.

As the full set of sensors is estimated, the sensors not used as measurements can 587 be compared with their reference simulation values to assess if the filter can correctly 588 estimate the full field of strains and accelerations by using only a limited set of sensors. 589 This evaluation is based on the Averaged Absolute Error (AAE) metric defined in [15] 590 as $\sum_{i=1}^{T} |y_{est,i} - y_{meas,i}|/T$ where $y_{est,i}$ and $y_{meas,i}$ are the estimated and measured 591 signals at the i-th time step, respectively, and T is the number of steps in the considered 592 time interval. This metric is compared with the Maximum Response Amplitude (MRA) 593 for each case in order to compare the average error with the maximum value assumed 594 by the signal. For this numerical application, the sensor estimation results are shown 595 only for the most complex case of parameter and input estimation, as all the others are 596 comparable. 597

The inputs used for this validation are a static load on the tip, a 20 Hz sinusoidal load (Figure 8) and a broadband continuous random signal (Figure 14). Data on the inputs is listed in Table 2.



Figure 14: Continuous random input

Table 2:	Load	case	data
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Signal type	Frequency	Amplitude
Static	$0~{\rm Hz}$	14.7 N
Sine	20 Hz	5 N
Random	0-500 Hz	14 N (max)

Each application case has been run for every possible load case. In some of the cases the static load has not been used since it makes the system unobservable, as shown in the theoretical observability analysis (section 4). White noise, with an amplitude in the order of $10^{-7}m/m$ and $10^{-3}m/s^2$ for strain and acceleration measurements respectively, is added to the generated numerical measurements. Both simulations and estimations use a time step of 0.1ms.

An overview of the performance of all the cases in terms of parameter estimation is shown in Table 3. This table shows the minimum, maximum and average relative error of the estimated parameters with respect to the reference values in the last 20% of the time interval, for which the estimated values have typically converged to stable values.
 Analysis of this interval gives an idea on the convergence to the correct value and how
 much the parameter varies in time after converging.

Unknown quantities	Load	H	Error [9	6]			
		Parameter	\min	avg	\max		
	static	\mathbf{E}	0.01	0.01	0.01		
E	sine	\mathbf{E}	0.01	0.02	0.02		
	random	E	-0.02	-0.02	-0.02		
	static	un	unobservable				
	sine	\mathbf{E}	-0.50	-0.22	0.00		
E, ρ		ρ	-0.30	-0.09	0.15		
	random	E	-0.23	-0.08	0.16		
		ρ	-0.29	-0.07	0.19		
	static	Е	-0.06	0.07	0.19		
degrading E	sine	E	-0.08	0.24	0.43		
	random	E	-0.04	0.16	0.38		
	static	unobservable					
E, input	sine	\mathbf{E}	-0.09	-0.08	-0.07		
	random	E	0.18	0.21	0.23		
	static	unobservable					
degrading E, input	sine	\mathbf{E}	0.07	0.39	0.75		
	random	E	0.27	0.56	0.97		

Table 3: Errors on estimated parameters in last 20% of time interval for numerical measurements

⁶¹³ 5.3.1. Constant Parameter Estimation

As a first application case, the Young's modulus is estimated. The density and the loads are known. The initial value of the augmented state related to the parameter is set to a value 25% higher than the reference value, and its covariance is calculated by setting $\delta_E = 10^{-3}Hz$. This parameter is observable for all considered load cases.

The evolution of the estimated parameter value over a 1s time interval is shown in Figure 15. For the static load case the estimated parameter shows a fast convergence to the reference value since the load is applied from the start. For the dynamic cases instead the value converges gradually to the correct one, since the initial input is zero and it gradually increases, allowing the filter to smoothly correct the value of E.



Figure 15: Estimated value of E with numerical data, for all load cases, with $\delta_E = 10^{-3} Hz$.

As a second application case, both Young's modulus and density are unknown. The results are shown in Figure 16. Only the dynamic load cases can be considered, since the density is unobservable for a static input. The initial errors on the augmented states related to E and ρ are respectively set at 25% and -30%. The observed convergence rate is slower and thus a larger δ_p of 1Hz is adopted. The converged parameter values estimated by the filter accurately match the reference ones.



Figure 16: Estimated values of E and ρ with numerical data, for sine and random loads, with $\delta_E = 1Hz$, $\delta_\rho = 1Hz$.

629 5.3.2. Degrading Parameter Estimation

A fundamental application of the presented methodology is the tracking of a parameter that deteriorates in time. In order to generate numerical data for the deterioration case, a simulation is run using the pROM during which the Young's modulus, starting from the correct value, decreases at a rate of 10%/s. It should be noticed how this type of simulation during which the parameters vary continuously is enabled by the use of the pROM, as it allows to update the model in an efficient way (i.e. without having to recreate the ROM).

⁶³⁷ The results of the estimation of only E, with $\delta_E = 1Hz$, are shown in Figure 17. The ⁶³⁸ large covariance value needed to allow the correct tracking of the parameter variation ⁶³⁹ causes measurement noise to propagate to the estimated parameter. This can be seen as ⁶⁴⁰ an acceptable solution as the average error remains small.



Figure 17: Estimated value of degrading E with numerical data, for all load cases, with $\delta_E = 1Hz$.

641 5.3.3. Parameter and Input Estimation

As a more complex application case, both the Young's modulus and the input are 642 considered as unknown. In this case the filter has to be able to discern between the 643 influence of a change of parameter values from a time-varying input. The results of the 644 coupled estimation of constant E ($\delta_E = 10^{-3}Hz$) and a dynamic load ($\omega_{u,2} = 80Hz$, 645 $\omega_{u,3} = 2000 Hz$) are shown in Figure 18. The system is unobservable for the static load 646 case. Both dynamic load cases show excellent results with a correct and fast estimation of 647 the parameters. The estimated input displays some noise given by the necessity for larger 648 related covariance values, as explained in subsection 5.2, to avoid lag. The estimated 649 sensor errors for both strain sensors and accelerometers are shown in Figure 19. It can 650 be observed how the AAE is always at least one order of magnitude lower that the MRA. 651 Some exceptions are in the accelerometers with lower MRA, due to accelerations being 652 more sensitive to the noise introduced by the input estimation. The estimation procedure 653 also shows good results for the case where the Young's modulus is varying in time, with 654 an unknown input and known density. The results are shown in Figure 20. Here, δ_E is 655 set to 1 Hz and thus more noise is added. 656



Figure 18: Estimated values of E and inputs with numerical data, for the sine and random load cases, with $\delta_E = 10^{-3}Hz$, $\omega_{u,2} = 80Hz$ and $\omega_{u,3} = 2000Hz$. The right figures show a zoom of the 0.9s to 1s range.



Figure 19: Estimated sensors error for the case shown in Figure 18.



Figure 20: Estimated value of degrading E and input with numerical data, for the sine and random load cases, with $\delta_E = 1Hz$, $\omega_{u,2} = 80Hz$ and $\omega_{u,3} = 2000Hz$. The right figures show a zoom of the 0.9s to 1s range.

⁶⁵⁷ 5.3.4. Parameter and input estimation with non-zero initial conditions

The former results have been obtained with measurements on a model starting at zero 658 initial conditions. In order to demonstrate the robustness of the filter, input and constant 659 Young's modulus are estimated based on a set of measurements with non-zero initial 660 conditions. This is obtained by simulating the system for 2s and discarding the interval 661 from the start up to 1s. The results of the coupled estimation, with the same setting as 662 in the former section, are shown in Figure 21. The estimated values show an excellent 663 convergence to the reference ones. The main difference with respect to former cases is 664 in the initial steps, where a less gradual convergence behavior is observed. This can be 665 explained by the fact that the filter is setup with different initial conditions, resulting in a 666



⁶⁶⁷ larger initial mismatch that is compensated with quick variations of estimated quantities.

Figure 21: Estimated values of E and inputs with numerical data and non-zero initial conditions, for the sine and random load cases, with $\delta_E = 10^{-3}Hz$, $\omega_{u,2} = 80Hz$ and $\omega_{u,3} = 2000Hz$. The right figures show a zoom of the 1.9s to 2s range.

668 6. Experimental Validation

In this section, the experimental setup (Figure 22) is presented and the estimation methodology is validated using experimental data.



Figure 22: Instrumented blade

The main parameters of the FE model (Figure 2a) representing the physical blade, 671 namely Young's modulus, Poisson's ratio, density and stiffness of the bolts, have been 672 identified by a modal testing procedure and model updating [42] using a commercial 673 FE pre- and post-processing software package [32]. The identified parameter values are 674 listed in Table 4. The damping parameters for the proportional Rayleigh model have 675 been identified using an optimization procedure aimed at finding the values that minimize 676 the error between the experimental and numerical FRFs for one input point (the shaker's 677 application point) and two representative outputs (accelerometers on sections D and H 678 in the global y direction). The identification has been carried out in the 0-50Hz range. 679

Table 4: Experimentally identified model parameters

$E \ [GPa]$	$ ho \; [Kg/m^3]$	ν	α	β	K [N/mm]
116.04	4873	0.34208	0.02	0.0005	33167.2

⁶⁸⁰ A zero-input acquisition has been done in order to measure the noise level for every ⁶⁸¹ sensor. The strain gages and accelerometers as used for data acquisition have noise levels ⁶⁸² in the order of $10^{-7}m/m$ and $10^{-3}m/s^2$, respectively.

The measurement campaign consisted of separately applying static and dynamic loads to the blade and acquiring data to be used for the estimation. A static load has been applied at the blade tip by suspending a mass of known value $m_s = 1.5 Kg$. A 20Hz sine load, following the one as used for the numerical validation, has been applied on a point at the middle of the bottom face of the blade by means of an electrodynamic modal shaker. The tip of the shaker was connected to the blade via a mechanical impedance sensor to measure the applied force.

The measurements have been acquired using a sampling period of 1ms. Since the estimator uses a time step of 0.1ms, these are linearly interpolated between sampling points in order to match the filter's time step. An overview of the relative errors for the parameter and input estimation using experimental data is shown in Table 5 for the different considered load cases. The observed errors are small for most cases, with larger variation ranges when compared with the numerical application cases. The different cases

are further discussed below. For each of them, the filter has been setup as described in⁶⁹⁶ subsection 5.2.

Table 5: Errors on estimated parameters in the last 20% of the considered time interval for experimental measurements

Unknown quantities	Load	$\begin{array}{c c} \mathbf{Error} \ [\%] \\ \mathbf{Parameter} & \mathbf{min} \ \mathbf{avg} \\ \mathbf{E} & \mid 0.03 \mid 0.04 \mid \end{array}$			
		Parameter	\min	avg	\max
	static	E	0.03	0.04	0.06
Ε	sine	E	1.95	2.00	2.05
	static	un	observa	ble	
E, ρ	ain e	E	2.09	3.45	5.01
	LoadError [%]ParameterminavgstaticE0.030.04sineE1.952.00static $unobservable$ $sine$ 2.093.45sine ρ -2.38-1.00static $unobservable$ $sine$ 0.46 0.52	-1.00	-0.08		
	static	un	observa	ble	
E, input	sine	E	0.46	0.52	0.63

698 6.1. Parameter Estimation

699 6.1.1. Young's modulus

For the case where the Young's modulus is estimated, with a known density and 700 input, the augmented state covariance value is calculated by choosing $\delta_E = 10^{-3} Hz$. 701 The observed error is close to zero for the static load input and 2% on average for 702 the sine load case. The results are shown in Figure 23. The values of the estimated 703 sensors as compared to the measured experimental values display a small error, as shown 704 in Figure 24. The main mismatch is present for the sensors mounted closer to the root 705 of the blade, where it is assumed that the boundary conditions have a larger influence on 706 the system dynamics and that the error can thus be attributed to the unmodeled system 707 dynamics related to these boundary conditions. Figure 24 includes also the AAE of 708 the simulated measurements generated through a forward simulation with the reference 709 710 parameter value. These show a similar behavior to estimated ones, proving that the error is not introduced by the filter but it is due to a mismatch in the model. 711



Figure 23: Estimated value of E with experimental data, for the static and sine load cases, with $\delta_E = 10^{-3} Hz$.



Figure 24: Estimated sensor errors for the case shown in Figure 23.

712 6.1.2. Young's Modulus and Density

For the case where the Young's modulus and density are jointly estimated, with a 713 known input, the augmented state covariance values are calculated by choosing $\delta_p = 1Hz$. 714 The static load case leads to an unobservable system, hence only the sine load case is 715 considered. The observed error has an average value of 1% for density and 3.45% for 716 stiffness, which can be considered acceptable. The results are shown in Figure 25. As 717 can be seen in Figure 26, the convergence of the ratio between the two parameters is 718 quicker then for the single values. A small error for the full field of sensors can also be 719 observed in this case, as shown in Figure 27. The larger required value for $\delta_p = 1Hz$ 720 introduces more noise as well as a low frequency variation of the estimated parameter 721 values. This could be attributed to a small model mismatch. This mismatch is assumed 722 to be filtered out by the smaller value of $\delta_E = 10^{-3} Hz$ in the former case where only the 723 Young's modulus is estimated. 724



Figure 25: Estimated values of E and ρ with experimental data, for the sine load case, with $\delta_E = 1Hz$, $\delta_\rho = 1Hz$.



Figure 26: Value of the ratio between estimated E and ρ for the case shown in Figure 25.



Figure 27: Estimated sensors error for the case shown in Figure 25.

725 6.2. Parameter and Input Estimation

For the joint estimation of the Young's modulus and an unknown input, the asso-726 ciated augmented state covariance values are calculated choosing $\delta_E = 10^{-3} Hz$ and 727 $\omega_{u,2} = 80Hz$. As the system is unobservable for a static load, the sine input load case 728 is considered. An average error on the estimated Young's modulus below 1% can be 729 observed, as shown in Figure 28. While the full field of strain sensors is accurately es-730 timated, the acceleration estimated sensors exhibit a larger error as compared to the 731 previous cases Figure 29. As explained above for the numerical case, this could be at-732 733 tributed to the high frequency load content introduced by the filter due to the larger augmented state covariance values. 734



Figure 28: Estimated values of E and input with experimental data, for the sine load case, with $\delta_E = 10^{-3}Hz$ and $\omega_{u,2} = 80Hz$.



Figure 29: Estimated sensors error for the case shown in Figure 28.

735 7. Conclusions

In this work, the integration of a parametric Reduced Order Model (pROM) in an 736 augmented Extended Kalman Filter (EKF) is presented, allowing for the joint estima-737 tion of states, inputs and material parameters of an industrial scale structural dynamic 738 system. A projection-based parametric Model Order Reduction (pMOR) scheme that 739 exploits the affine dependency on the parameters of the full model and uses a constant 740 global reduction basis is proposed to achieve an efficient reduced model that maintains 741 an explicit dependency on the parameters. The integration of the pROM in the EKF 742 algorithm enables the efficient parameter estimation. 743

An observability analysis shows that in a general case a number of position-level
measurements at least equal to the number of unknown inputs and parameters is required. Furthermore, several considerations on the observability of specific combinations
of unknown mass and stiffness parameters and inputs are discussed.

The proposed methodology is applied to a scaled wind turbine blade. The pROM 748 is numerically validated by comparing time and frequency results with a standard (i.e. 749 non-parametric) Reduced Order Model (ROM). Estimation of parameters, inputs and 750 751 states on both numerical and experimental data shows that the proposed approach is able to correctly estimate the unknown quantities together with the full field of strains 752 and accelerations using a small set of strain sensors. Practical considerations on the 753 setup of the filter are discussed, with a focus on the augmented state covariance values; 754 these show how using an adapting scheme for the selection of covariance values would be 755 helpful and should be explored in future research. 756

Future efforts will be aimed at estimating anisotropic constitutive material parameters, as well as developing an extended Optimal Sensor Placement (OSP) approach that considers the augmented states related to the material parameters.

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